

Non-overlapping descents and ascents in stack-sortable permutations

Sergey Kitaev* and Philip B. Zhang†

October 26, 2023

Abstract. The Eulerian polynomials $A_n(x)$ give the distribution of descents over permutations. It is also known that the distribution of descents over stack-sortable permutations (i.e. permutations sortable by a certain algorithm whose internal storage is limited to a single stack data structure) is given by the Narayana numbers $\frac{1}{n} \binom{n}{k} \binom{n}{k+1}$. On the other hand, as a corollary of a much more general result, the distribution of the statistic “maximum number of non-overlapping descents”, MND, over all permutations is given by $\sum_{n,k \geq 0} D_{n,k} x^k \frac{t^n}{n!} = \frac{e^t}{1-x(1+(t-1)e^t)}$.

In this paper, we show that the distribution of MND over stack-sortable permutations is given by $\frac{1}{n+1} \binom{n+1}{2k+1} \binom{n+k}{k}$. We give two proofs of the result via bijections with rooted plane (binary) trees allowing us to control MND. Moreover, we show combinatorially that MND is equidistributed with the statistic MNA, the maximum number of non-overlapping ascents, over stack-sortable permutations. The last fact is obtained by establishing an involution on stack-sortable permutations that gives equidistribution of 8 statistics.

AMS Classification 2010: 05A15

Keywords: stack-sortable permutation; descent; ascent; pattern-avoidance

1 Introduction

A permutation of length n is a rearrangement of the set $[n] := \{1, 2, \dots, n\}$. Denote by S_n the set of permutations of $[n]$ and let ε be the empty permu-

*Department of Mathematics and Statistics, University of Strathclyde, 26 Richmond Street, Glasgow G1 1XH, United Kingdom. **Email:** sergey.kitaev@strath.ac.uk.

†College of Mathematical Science, Tianjin Normal University, Tianjin 300387, P. R. China **Email:** zhang@tjnu.edu.cn.

tation. For a permutation $\pi = \pi_1\pi_2\cdots\pi_n$ with $\pi_i = n$, the *stack-sorting operator* \mathcal{S} is defined recursively as follows, where $\mathcal{S}(\varepsilon) = \varepsilon$,

$$\mathcal{S}(\pi) = \mathcal{S}(\pi_1\cdots\pi_{i-1})\mathcal{S}(\pi_{i+1}\cdots\pi_n)n.$$

A permutation π is *stack-sortable* (S-S) if $\mathcal{S}(\pi) = 12\cdots n$. Let SS_n denote the set of S-S permutations of length n . The stack-sorting operator appears in numerous studies in the mathematics and theoretical computer science literature (e.g. see [2] and references therein).

The permutations in SS_n are counted by the Catalan numbers $C_n = \frac{1}{n+1}\binom{2n}{n}$, and they are precisely the set of 231-avoiding permutations, where a permutation $\pi_1\pi_2\cdots\pi_n$ avoids a pattern $p = p_1p_2\cdots p_k$ (which is also a permutation) if there is no subsequence $\pi_{i_1}\pi_{i_2}\cdots\pi_{i_k}$ such that $\pi_{i_j} < \pi_{i_m}$ if and only if $p_j < p_m$ [5, 7]. Let $S_n(p)$ denote the set of p -avoiding permutations of length n . Then $SS_n = S_n(231)$. S-S permutations have a nice recursive structure: if $\pi = AnB \in S_n(231)$ then $A < B$ (i.e. every element in A is less than any element in B) and A and B are (possibly empty) 231-avoiding permutations independent from each other.

For a permutation $\pi = \pi_1\pi_2\cdots\pi_n$, the *descent* (resp., *ascent*) *statistic* on π , $\text{des}(\pi)$ (resp., $\text{asc}(\pi)$), is defined as the number of $i \in [n-1]$ such that $\pi_i > \pi_{i+1}$ (resp., $\pi_i < \pi_{i+1}$). For example, $\text{des}(562413)=2$ and $\text{asc}(35124)=3$. The distribution of descents (or ascents) over S_n is given by the *Eulerian polynomial*

$$A_n(x) := \sum_{\pi \in S_n} x^{\text{des}(\pi)} = \sum_{k=1}^n k!S(n, k)(x-1)^{n-k}$$

where $S(n, k)$ is the *Stirling numbers of the second kind*. On the other hand, the distribution of descents over SS_n (i.e. 231-avoiding permutations) is given by the Narayana numbers

$$\frac{1}{n}\binom{n}{k}\binom{n}{k+1}$$

(see [9]). To complete the picture, the distribution of descents over 123-avoiding permutations is given by the following formula [1, 3], where t and x correspond to the length and the number of descents:

$$\frac{-1 + 2tx + 2t^2x - 2tx^2 - 4t^2x^2 + 2t^2x^3 + \sqrt{1 - 4tx - 4t^2x + 4t^2x^2}}{2tx^2(tx - 1 - t)}.$$

The distribution for the number of 321-avoiding permutations of length n with k descents [8, A091156] is

$$\frac{1}{n+1} \binom{n+1}{k} \sum_{j=0}^{n-2k} \binom{k+j-1}{k-1} \binom{n+1-k}{n-2k-j}. \quad (1)$$

The notion of the *maximum number of non-overlapping occurrences* of a *consecutive pattern* (that is, occurrences of a pattern defined above with the requirement of $i_2 = i_1 + 1$, $i_3 = i_2 + 1$, etc) in a permutation has been considered in [6]. It turns out that the distribution of the maximum number of non-overlapping occurrences of a consecutive pattern can be expressed in terms of the *exponential generating function* (e.g.f.) for the number of permutations *avoiding* the pattern. In particular, since the e.g.f. for permutations with no descents is clearly e^t , we have [6] that

$$\sum_{n,k \geq 0} D_{n,k} x^k \frac{t^n}{n!} = \frac{e^t}{1 - x(1 + (t-1)e^t)}$$

where $D_{n,k}$ is the number of permutations with k non-overlapping descents (but without $k+1$ non-overlapping descents).

The main focus of this paper is the study of the *maximum number of non-overlapping descents* (MND) and *non-overlapping ascents* (MNA) over S-S permutations. For example, $\text{MND}(13254) = \text{MND}(32154) = 2$ while $2 = \text{des}(13254) \neq \text{des}(32154) = 3$. In Section 3 we will provide an involution on S-S permutations that preserves 8 statistics and shows that the distribution of MNA is the same as that of MND. Further, we will prove in two ways (in Sections 4 and 5), via establishing appropriate bijections with rooted plane (binary) trees, that the distribution of MND over S-S (i.e. 231-avoidable permutations) is given by

$$\frac{1}{n+1} \binom{n+1}{2k+1} \binom{n+k}{k}. \quad (2)$$

The respective numbers are recorded in [8, A108759] and the formula is derived in [4]. Finally, in Section 6 we suggest directions for further research.

2 Preliminaries

In this section we provide necessary basic definitions and notation used in the paper.

2.1 Trees.

A *rooted plane tree* or an *ordered tree*, consists of a set of vertices each of which has a (possibly empty) linearly ordered list of vertices associated with it called its *children*. One of the vertices of the tree is called the *root*. A vertex with no children is a *leaf*. A vertex in a rooted plane tree is *internal* if it is neither a leaf nor the root. Hence, the vertices are partitioned into three classes: root, internal vertices, leaves. Rooted plane trees can be produced as follows:

- A single vertex with no children is a rooted plane tree. That vertex is the root.
- If T_1, \dots, T_k is an ordered list of rooted plane trees with roots r_1, \dots, r_k and no vertices in common, then a rooted plane tree T can be constructed by choosing an unused vertex r to be the root, letting its i -th child be r_i (r_1, \dots, r_k are not roots any more).

An internal vertex in a tree is *marked* if it has a leaf as its child. The formula (2) gives the number of rooted plane trees with n edges and k marked vertices [4]. In this paper, we call these objects *marked trees*. The same formula counts *full binary trees* (ordered trees where there are no or two children for each vertex) on $2n$ edges by the value k of the following statistic X described in [8, A108759]. Delete all right edges in a given full binary tree T with $2n$ edges leaving the left edges in place. This partitions the left edges into line segments (or paths) of lengths say $\ell_1, \ell_2, \dots, \ell_t$, with $\sum_{i=1}^t \ell_i = n$. Then $X(T) = \sum_{i=1}^t \lfloor \frac{\ell_i}{2} \rfloor$. This result is implicit in [10]. For example, $X(T) = 2$ for T in Figure 1. Note that an interpretation of (2) on *Dyck paths* can be given via a standard bijection with full binary trees [8, A108759], but Dyck paths are not to be considered in this paper.

2.2 Permutations.

Section 1 introduces the notion of a pattern-avoiding permutation, the structure of S-S permutations (i.e. 231-avoiding permutations) and the statistics asc, des, MNA and MND. In this paper, we also need the notions of statistics from [5] introduced in Table 1. For example, $\text{lmin}(52341) = \text{rmax}(52341) = \text{rmin}(413625) = 3$, $\text{ldr}(526341) = \text{rar}(413625) = 2$. Moreover, the *reverse* of $\pi = \pi_1 \pi_2 \dots \pi_n$ is the permutation $r(\pi) = \pi_n \pi_{n-1} \dots \pi_1$ and the *complement* of π is the permutation $c(\pi)$ obtained from π by replacing each π_i by $n + 1 - \pi_i$. For example, $r(31425) = 52413$ and $c(31425) = 35241$. For a sequence

Stat	Definition on a permutation $\pi_1\pi_2\cdots\pi_n$
lmin	number of left-to-right minima = $ \{i \mid \pi_j > \pi_i \text{ for all } j < i\} $
rmin	number of right-to-left minima = $ \{i \mid \pi_i < \pi_j \text{ for all } j > i\} $
ldr	length of leftmost decreasing run = $\max\{i \mid \pi_1 > \pi_2 > \cdots > \pi_i\}$
rar	length of rightmost ascending run = $\max\{i \geq 1 \mid \pi_n > \cdots > \pi_{n-i+1}\}$
rmax	number of right-to-left maxima = $ \{i \mid \pi_i > \pi_j \text{ for all } j > i\} $

Table 1: Permutation statistics in this paper.

of distinct numbers $x = x_1x_2\cdots x_k$, the *reduced form* of x , $\text{red}(x)$, is obtained from x by replacing the i -th smallest number by i . For example, $\text{red}(2547)=1324$.

k -tuples of (permutation) statistics (s_1, s_2, \dots, s_k) and $(s'_1, s'_2, \dots, s'_k)$ are equidistributed over a set S if

$$\sum_{a \in S} t_1^{s_1(a)} t_2^{s_2(a)} \cdots t_k^{s_k(a)} = \sum_{a \in S} t_1^{s'_1(a)} t_2^{s'_2(a)} \cdots t_k^{s'_k(a)}.$$

For example, (asc, des) and (des, asc) are equidistributed over S_n for any $n \geq 0$, and this fact is trivial from applying the reverse (or complement) to all permutations in S_n . The fact that (MNA, MND) is equidistributed with (MND, MNA) over S-S permutations is not trivial, and it will follow from the more general equidistribution result (involving 8 statistics) in Section 3.

3 Equidistribution of MND and MNA over S-S permutations

The equidistribution of MND and MNA over S-S permutations follows from the following more general result.

Theorem 1. *The following 8-tuples of statistics are equidistributed over $S_n(231)$ for $n \geq 0$:*

$$(MND, MNA, asc, des, ldr, rar, lmin, rmin)$$

$$(MNA, MND, des, asc, rar, ldr, rmin, lmin).$$

Proof. Recall that if $\pi = AnB \in S_n(231)$ then $A < B$. We define the following recursive map f on $S_n(231)$ that is easy to check by induction on n to be an involution (that is, f^2 is the identity map):

- $f(\varepsilon) = \varepsilon$.
- For $\pi = AnB \in S_n(231)$, $n \geq 1$, where A and B are possibly empty, $f(\pi) = f(\text{red}(B))n(f(A))^+$ where $(f(A))^+$ is formed by the largest elements in $f(\pi)$ excluding n and $\text{red}((f(A))^+) = f(A)$. That is, f recursively swaps the largest and smallest elements in π below the largest element. It is easy to see by induction on n that $f(\pi) \in S_n(231)$.

We prove by induction on n that the map f respects the statistics. The base cases of $n \in \{0, 1\}$ are trivial, so assume $n \geq 2$. We now apply the inductive hypothesis in the following three cases (some of which, but not all, result in the same derivations):

$$\boxed{A \neq \varepsilon, B \neq \varepsilon.}$$

$$\begin{aligned} \text{ldr}(\pi) &= \text{ldr}(A) = \text{rar}(f(A)) = \text{rar}(f(\pi)). \\ \text{rar}(\pi) &= \text{rar}(\text{red}(B)) = \text{ldr}(f(\text{red}(B))) = \text{ldr}(f(\pi)). \\ \text{lmin}(\pi) &= \text{lmin}(A) = \text{rmin}(f(A)) = \text{rmin}(f(\pi)). \\ \text{rmin}(\pi) &= \text{rmin}(\text{red}(B)) = \text{lmin}(f(\text{red}(B))) = \text{lmin}(f(\pi)). \\ \text{asc}(\pi) &= \text{asc}(A) + 1 + \text{asc}(\text{red}(B)) = \text{des}(f(A)) + 1 + \text{des}(f(\text{red}(B))) = \text{des}(f(\pi)). \\ \text{des}(\pi) &= \text{des}(A) + 1 + \text{des}(\text{red}(B)) = \text{asc}(f(A)) + 1 + \text{asc}(f(\text{red}(B))) = \text{asc}(f(\pi)). \end{aligned}$$

Now, suppose that $\text{ldr}(\text{red}(B))$ is even. Then, because $\text{ldr}(\pi) = \text{rar}(f(\pi))$, the element n does not contribute to $\text{MND}(\pi)$ and $\text{MNA}(f(\pi))$, so we have

$$\begin{aligned} \text{MND}(\pi) &= \text{MND}(A) + \text{MND}(\text{red}(B)) = \text{MNA}(f(A)) + \text{MNA}(f(\text{red}(B))) \\ &= \text{MNA}(f(\pi)). \end{aligned}$$

On the other hand, if $\text{ldr}(\text{red}(B))$ is odd, then n contributes one extra non-overlapping descent in π and one extra non-overlapping ascent in $f(\pi)$:

$$\begin{aligned} \text{MND}(\pi) &= \text{MND}(A) + 1 + \text{MND}(\text{red}(B)) = \text{MNA}(f(A)) + 1 + \text{MNA}(f(\text{red}(B))) \\ &= \text{MNA}(f(\pi)). \end{aligned}$$

Similarly, suppose that $\text{rar}(A)$ is even. Then, because $\text{rar}(\pi) = \text{ldr}(f(\pi))$, the element n does not contribute to $\text{MNA}(\pi)$ and $\text{MND}(f(\pi))$, so we have

$$\begin{aligned} \text{MNA}(\pi) &= \text{MNA}(A) + \text{MNA}(\text{red}(B)) = \text{MND}(f(A)) + \text{MND}(f(\text{red}(B))) \\ &= \text{MND}(f(\pi)). \end{aligned}$$

Finally, if $\text{rar}(A)$ is odd, then n contributes one extra non-overlapping ascent in π and one extra non-overlapping descent in $f(\pi)$:

$$\begin{aligned} \text{MNA}(\pi) &= \text{MNA}(A) + 1 + \text{MNA}(\text{red}(B)) = \text{MND}(f(A)) + 1 + \text{MND}(f(\text{red}(B))) \\ &= \text{MND}(f(\pi)). \end{aligned}$$

$$\boxed{A = \varepsilon, B \neq \varepsilon.}$$

$$\begin{aligned} \text{ldr}(\pi) &= 1 + \text{ldr}(\text{red}(B)) = 1 + \text{rar}(f(\text{red}(B))) = \text{rar}(f(\pi)). \\ \text{rar}(\pi) &= \text{rar}(\text{red}(B)) = \text{ldr}(f(\text{red}(B))) = \text{ldr}(f(\pi)). \\ \text{lmin}(\pi) &= 1 + \text{lmin}(\text{red}(B)) = 1 + \text{rmin}(f(\text{red}(B))) = \text{rmin}(f(\pi)). \\ \text{rmin}(\pi) &= \text{rmin}(\text{red}(B)) = \text{lmin}(f(\text{red}(B))) = \text{lmin}(f(\pi)). \end{aligned}$$

$$\begin{aligned}\text{asc}(\pi) &= \text{asc}(\text{red}(B)) = \text{des}(f(\text{red}(B))) = \text{des}(f(\pi)). \\ \text{des}(\pi) &= 1 + \text{des}(\text{red}(B)) = 1 + \text{asc}(f(\text{red}(B))) = \text{asc}(f(\pi)).\end{aligned}$$

Now, suppose that $\text{ldr}(\text{red}(B))$ is even. Then, because $\text{ldr}(\pi) = \text{rar}(f(\pi))$, the element n does not contribute to $\text{MND}(\pi)$ and $\text{MNA}(f(\pi))$, so we have $\text{MND}(\pi) = \text{MND}(\text{red}(B)) = \text{MNA}(f(\text{red}(B))) = \text{MNA}(f(\pi))$.

However, if $\text{ldr}(\text{red}(B))$ is odd, then n contributes one extra non-overlapping descent in π and one extra non-overlapping ascent in $f(\pi)$:

$$\text{MND}(\pi) = 1 + \text{MND}(\text{red}(B)) = 1 + \text{MNA}(f(\text{red}(B))) = \text{MNA}(f(\pi)).$$

$$\text{Finally, } \text{MNA}(\pi) = \text{MNA}(\text{red}(B)) = \text{MND}(f(\text{red}(B))) = \text{MND}(f(\pi)).$$

$$\boxed{A \neq \varepsilon, B = \varepsilon.}$$

$$\text{ldr}(\pi) = \text{ldr}(A) = \text{rar}(f(A)) = \text{rar}(f(\pi)).$$

$$\text{rar}(\pi) = 1 + \text{rar}(A) = 1 + \text{ldr}(f(A)) = \text{ldr}(f(\pi)).$$

$$\text{lmin}(\pi) = \text{lmin}(A) = \text{rmin}(f(A)) = \text{rmin}(f(\pi)).$$

$$\text{rmin}(\pi) = 1 + \text{rmin}(A) = 1 + \text{lmin}(f(A)) = \text{lmin}(f(\pi)).$$

$$\text{asc}(\pi) = \text{asc}(A) + 1 = \text{des}(f(A)) + 1 = \text{des}(f(\pi)).$$

$$\text{des}(\pi) = \text{des}(A) = \text{asc}(f(A)) = \text{asc}(f(\pi)).$$

$$\text{Now, } \text{MND}(\pi) = \text{MND}(A) = \text{MNA}(f(A)) = \text{MNA}(f(\pi)).$$

Further, suppose that $\text{rar}(A)$ is even. Then, because $\text{rar}(\pi) = \text{ldr}(f(\pi))$, the element n does not contribute to $\text{MNA}(\pi)$ and $\text{MND}(f(\pi))$, so we have $\text{MNA}(\pi) = \text{MNA}(A) = \text{MND}(f(A)) = \text{MND}(f(\pi))$.

Finally, if $\text{rar}(A)$ is odd, then n contributes one extra non-overlapping ascent in π and one extra non-overlapping descent in $f(\pi)$:

$$\text{MNA}(\pi) = \text{MNA}(A) + 1 = \text{MND}(f(A)) + 1 = \text{MND}(f(\pi)).$$

This completes our proof. \square

4 MND on S-S permutations and binary trees

To achieve the desired result, we need the following theorem, the proof of which introduces the bijection g on S-S permutations.

Theorem 2. *The statistics ldr and rmax are equidistributed on S-S permutations.*

Proof. It is easy to see using the fact that S-S permutations are precisely 231-avoiding permutations that the structure of any S-S permutation is $x_\ell x_{\ell-1} \cdots x_2 x_1 A_1 A_2 \cdots A_{\ell-1} A_\ell$ where

- x_i 's are left-to-right minima (in particular, $x_1 = 1$);
- $1 = x_1 < A_1 < x_2 < A_2 < \cdots < x_\ell < A_\ell$ (i.e., the elements in A_i are a permutation of the set $\{x_i + 1, x_i + 2, \dots, x_{i+1} - 1\}$ assuming $x_{\ell+1} := n + 1$);

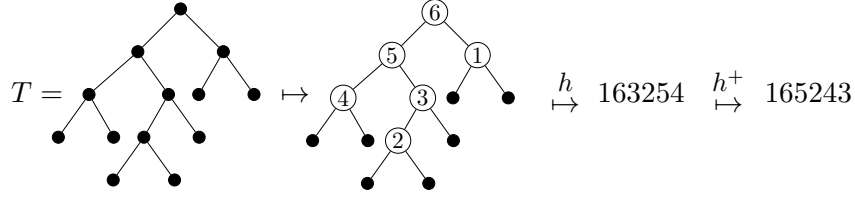


Figure 1: An illustration of the steps in the proof of Theorem 3. Note that $X(T) = \lfloor \frac{3}{2} \rfloor + \lfloor \frac{2}{2} \rfloor + \lfloor \frac{1}{2} \rfloor = 2 = \text{MND}(165243)$.

- each A_i is a (possibly empty) 231-avoiding permutation.

On the other hand, to avoid the pattern 231, the structure of any S-S permutation is $B_1 n B_2(n-1) \cdots B_\ell(n-\ell+1)$ where

- $n(n-1) \cdots (n-\ell+1)$ is the sequence of right-to-left maxima formed by the largest elements;
- $B_1 < B_2 < \cdots < B_\ell$;
- each B_i is a (possibly empty) 231-avoiding permutation.

The map g takes an S-S permutation π with $\text{rmax}(\pi) = \ell$ and the structure described above and sends it to the S-S permutation $g(\pi)$ with $\text{ldr}(g(\pi)) = \ell$ and the structure described above so that $\text{red}(A_i) = \text{red}(B_i)$ for $i = 1, 2, \dots, \ell$. That is straightforward to see that g is injective and surjective and hence bijective proving the equidistribution of ldr and rmax . \square

The fact that the distribution of MND on S-S permutations is given by (2) follows from the next theorem (recall the definition of the statistic X in Section 2.1). We refer to Figure 1 for an illustration of the steps in the proof of Theorem 3.

Theorem 3. *The set $\{T \mid T \text{ is a full binary tree with } 2n \text{ edges; } X(T) = k\}$ is in one-to-one correspondence with the set $\{\pi \mid \pi \text{ is an S-S permutation of length } n; \text{MND}(\pi) = k\}$.*

Proof. Label the non-leaf vertices in a full binary tree T with $2n$ edges using the pre-order traversal (that is, visiting the leftmost yet unvisited vertices first) and decreasing order of labels starting from n . Then $T = T_L n T_R$ where n is the root, T_L is the left subtree and T_R is the right subtree. Now, define recursively the map h from labeled full binary trees to 231-avoiding

permutations by $h(T) := h(T_R)nh(T_L)$, and as the base case, h maps the one-vertex tree to ε . In particular, if $h(T_L) \neq \varepsilon$, the permutation $h(T_L)$ is formed by the largest elements below n and hence the outcome is a 231-avoiding permutation. Note that if $h(T) = AnB$, then $h^+(T) := Ang(B)$ where g is defined in the proof of Theorem 2 and by $g(B)$ we actually mean first computing $B' = g(\text{red}(B))$ and then taking the order-isomorphic to B' permutation formed by the largest elements below n . Clearly, h^+ is a bijection. Also, note that $\text{lpath}(T) = \text{rmax}(h(T))$, where $\text{lpath}(T)$ is the length of the leftmost path in T , that is, the number of edges on the path from the root to the leftmost leaf.

Next, we prove by induction on n , with the obvious base cases of $n = 0, 1$, that if $X(T) = k$ then $\text{MND}(h^+(T)) = k$. We consider two cases given by the parity of $\text{lpath}(T)$. Suppose that $\text{lpath}(T)$ is odd. Then, by definition of X , $X(T) = X(T_L n T_R) =$

$$X(T_L) + X(T_R) = \text{MND}(h^+(T_L)) + \text{MND}(h^+(T_R)) = \text{MND}(h^+(T))$$

where the last equality follows from the fact that $\text{rmax}(h(T_L)) = \text{ldr}(h^+(T_L))$ is even and the element n in $h^+(T)$ does not increase MND . On the other hand, if $\text{lpath}(T)$ is even then $X(T) = X(T_L n T_R) =$

$$X(T_L) + 1 + X(T_R) = \text{MND}(h^+(T_L)) + 1 + \text{MND}(h^+(T_R)) = \text{MND}(h^+(T))$$

where the last equality follows from the fact that $\text{rmax}(h(T_L)) = \text{ldr}(h^+(T_L))$ is odd and the element n in $h^+(T)$, together with the leftmost decreasing run in $h^+(T_L)$, creates an extra occurrence of a non-overlapping descent in $h^+(T)$, and the theorem is proved. \square

5 MND on S-S permutations and marked trees

This section gives an alternative proof (to the one presented in Section 4) of the fact that the distribution of MND on S-S permutations is given by (2), and we present it in the next theorem. (Recall relevant definitions in Section 2.1). The bijection in the proof of Theorem 4 is the most involved out of the three bijections in this paper.

Theorem 4. *The set*

$$\{T \mid T \text{ is a rooted plane tree with } n \text{ edges and } k \text{ marked vertices}\}$$

is in one-to-one correspondence with the set

$$\{\pi \mid \pi \text{ is an S-S permutation of length } n; \text{MND}(\pi) = k\}.$$

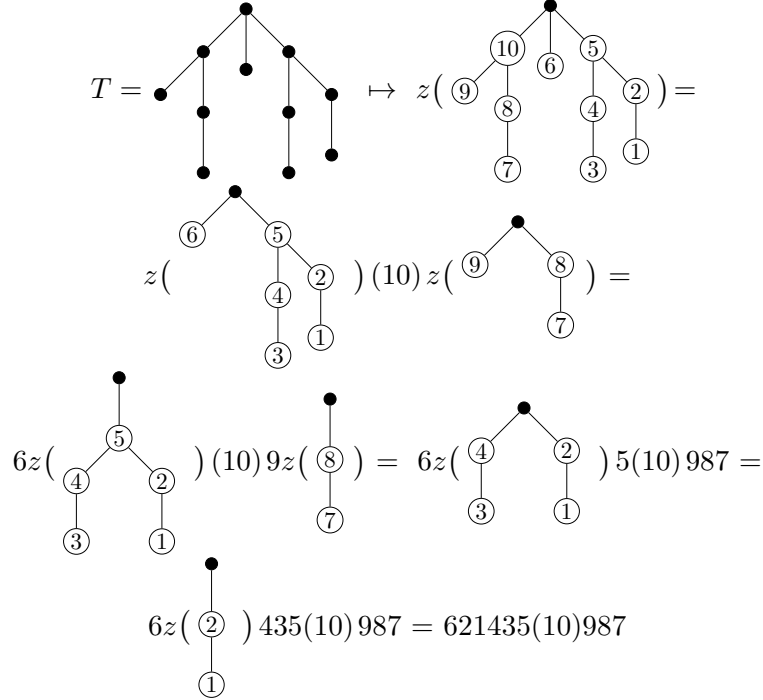


Figure 2: Computing $z(T)$ in the proof of Theorem 4 that results in the permutation $\pi = 621435(10)987$. Note that $\text{ldr}(\pi)$ is odd, which is consistent with T 's root having a leaf as a child, and $M(T) = \text{MND}(\pi) = 4$ as the marked vertices in T are 2, 4, 8, 10.

Proof. Let $M(T)$ denote the number of marked vertices in a rooted plane tree T .

We define the map z from the set of trees to the set of permutations as follows. The one vertex tree (with no edges, $n = 0$) is mapped by z to the empty permutation ε . For the rest of the recursive description of z it is convenient to label all non-root vertices in each rooted plane tree using the pre-order traversal (beginning from the root and going in the leftmost available direction) and starting assigning the largest label n (to the leftmost child of the root) and then assigning the labels in the descending order (similarly to our labelling in the proof of Theorem 3). To define the image $z(T)$ for a tree T in the set with $n \geq 1$, we distinguish the following five cases in which we check by induction on n , with the trivial base case of $n = 0$, that

- $\text{ldr}(z(T))$ is even if the root of T has no leaf as a child and $\text{ldr}(z(T))$ is odd otherwise, and
- $M(T) = \text{MND}(z(T))$.

We refer to Figure 2 for an example involving all cases. So, the cases are:

- (1) T has a single child r_1 labeled by n that is the root of the subtree A and no children of r_1 is a leaf. Let $z(T) := z(A)n$. Note that $\text{ldr}(z(A))$ is even by the inductive hypothesis, and $\text{ldr}(z(T))$ is even too. Also, $\text{MND}(z(T)) = \text{MND}(z(A)) = M(A) = M(T)$ since r_1 is not marked in T .
- (2) T has a single child r_1 labeled by n that is the root of the subtree A and at least one child of r_1 is a leaf, so r_1 is marked. Let $z(T) := nz(A)$. Note that $\text{ldr}(z(A))$ is odd by the inductive hypothesis, and hence $\text{ldr}(z(T))$ is even. Then the element n in $z(T)$ contributes an extra non-overlapping descent, which is consistent with r_1 being marked: $\text{MND}(z(T)) = 1 + \text{MND}(z(A)) = 1 + M(A) = M(T)$.
- (3) The leftmost child r_1 (labelled by n) of the root is a leaf, the root has at least one more child, but none of the other children of the root is a leaf. Let the subtree A be T without r_1 . Then $z(T) := nz(A)$. Note that $\text{ldr}(z(A))$ is even by the inductive hypothesis, and hence $\text{ldr}(z(T))$ is odd (which makes the outcome here be different from that in case (2)). Clearly, $\text{MND}(z(T)) = \text{MND}(z(A)) = M(A) = M(T)$.
- (4) The root has no leaves, the leftmost child r_1 (labelled by n) of the root is the root of the subtree B (with at least one edge), and the rest of T is the subtree A with at least one edge (and root having no leaves as children). Let $z(T) := z(A)nZ(B)$ where $Z(B)$ is formed by the largest elements below n . Note that $\text{ldr}(z(T)) = \text{ldr}(z(A))$ is even. Furthermore,
 - if r_1 has no leaf among its children, then $\text{ldr}(z(B))$ is even and $\text{MND}(z(T)) = \text{MND}(z(A)) + \text{MND}(z(B)) = M(A) + M(B) = M(T)$;
 - if r_1 has a leaf as a child, then $\text{ldr}(z(B))$ is odd, r_1 is marked and $\text{MND}(z(T)) = \text{MND}(z(A)) + 1 + \text{MND}(z(B)) = M(A) + 1 + M(B) = M(T)$.

We remark that it is essential for our goals to consider cases (3) and (4) separately rather than allowing B be the one-vertex tree in case (4).

- (5) The root has a leaf, the leftmost child r_1 (labelled by n) of the root is the root of the possibly one-vertex subtree B , and the rest of T is

the subtree A with at least one edge and root having a leaf as a child. Let $z(T) := z(A)nZ(B)$ where $Z(B)$ is formed by the largest elements below n . Note that $\text{ldr}(z(T)) = \text{ldr}(z(A))$ is odd. Furthermore,

- if r_1 has no leaf among its children, then $\text{ldr}(z(B))$ is even (possibly 0) and $\text{MND}(z(T)) = \text{MND}(z(A)) + \text{MND}(z(B)) = \text{M}(A) + \text{M}(B) = \text{M}(T)$;
- if r_1 has a leaf as a child, then $\text{ldr}(z(B))$ is odd, r_1 is marked and $\text{MND}(z(T)) = \text{MND}(z(A)) + 1 + \text{MND}(z(B)) = \text{M}(A) + 1 + \text{M}(B) = \text{M}(T)$.

Note that if B is the one-vertex tree in case (5), this case is different from case (3).

It is straightforward to see that the cases above are disjoint and any given rooted plane tree belongs to one of the cases. In particular, if the root of a tree has a single child then we are in case (1) or case (2). Moreover, the map z is clearly injective and surjective (as every 231-avoiding permutation appears exactly once as the image which is easy to see by considering where n is placed) and hence z is a bijection giving the desired result. \square

6 Directions for further research

In this paper, we find the distribution of the maximum number of non-overlapping descents and ascents over S-S permutations, which are precisely 231-avoiding permutations. Using trivial bijections on permutations (reverse, complement, and their composition), our results provide the distribution of MND and MNA (given by (2)) over p -avoiding permutations, where $p \in \{132, 213, 231, 312\}$. Additionally note that MND on 321-avoiding permutations (equivalently, MNA on 123-avoiding permutations) is given by (1), because 321-avoiding permutations do not have occurrences of overlapping descents. However, finding distribution of MNA on 321-avoiding permutations is still an open problem (see Table 2 for the respective distribution).

Open directions for research include finding distributions of MND and MNA over other (pattern-avoiding) classes of permutations, in particular, over 2-stack sortable permutations (permutations sortable by two applications of the stack-sorting operator \mathcal{S}).

Acknowledgments. The first author is grateful to Tianjin Normal University for its hospitality. The work of the second author was supported by the National Science Foundation of China (No. 12171362).

$n \backslash k$	0	1	2	3	4	5	6
1	1						
2	1	1					
3	0	5					
4	0	8	6				
5	0	5	37				
6	0	0	89	43			
7	0	0	98	331			
8	0	0	42	1036	352		
9	0	0	0	1644	3218		
10	0	0	0	1320	12362	3114	
11	0	0	0	429	25498	32859	
12	0	0	0	0	29744	149264	29004

Table 2: Distribution of MNA (resp., MND) on 321-avoiding (resp., 123-avoiding) permutations, where n is the length of permutations and k is the number of occurrences of the statistic.

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